

第7講 差分の差による政策評価

パネル調査において、介入後の「処置群と対照群の差」と介入前の「処置群と対照群の差」の差を指す。¹

- 「処置群での介入後と介入前の差」と「対象群での介入後と介入前の差」の差に等しい。
 - 介入プログラムの効果+誤差。パネルデータ分析における固定効果推定と同じ。

y_a : 政策が実施される前の結果変数

y_{1b} : 「もし政策の対象となる集団に所属した場合における」政策実施後の結果変数

y_{1b} : 「もし政策の対象となる集団に所属しない場合における」政策実施後の結果

t 時点でのグループ所属インディケーター z_t

$$y_t = zy_{1b} + (1 - z)y_{0b}$$

δ を b 時点での測定値であれば $\delta = 1$ 、 a 時点の測定値であれば、 $\delta = 0$ とするインディケータとすると、

$$y = \delta y_b + (1 - \delta) y_a = \delta \{zy_{1b} + (1 - z)y_{0b}\} + (1 - \delta) y_a$$

$$\begin{aligned} DID &= \{E(y | z = 1, \delta = 1) - E(y | z = 1, \delta = 0)\} - \{E(y | z = 0, \delta = 1) - E(y | z = 0, \delta = 0)\} \\ &= E(y_{1b} - y_a | z = 1) - E(y_{0b} - y_a | z = 0) \end{aligned}$$

処置群での因果効果 (TET)

$$TET = E(y_{1b} - y_{0b} | z = 1)$$

$E(y_{0b} | z = 1)$ は観察できない。

「もし政策の対象とならなかったときの経時変化が 2 つのグループ間で等しい」

$$\Rightarrow E(y_{0b} - y_a | z = 1) = E(y_{0b} - y_a | z = 0)$$

ならば

$$DID = E(y_{1b} - y_a | z = 1) - E(y_{0b} - y_a | z = 1) + \{E(y_{0b} - y_a | z = 1) - E(y_{0b} - y_a | z = 0)\}$$

¹以下は、星野 (2009, pp.103-111) からの引用である。

$$= E(y_{1b} - y_a | z = 1) - E(y_{0b} - y_a | z = 1)$$

$$= E(y_{1b} - y_a | z = 1)$$

$$= TET$$

何も介入しなければ結果変数が時間的な変化を起こさない場合（成長なし）

$y_{0b} = y_a$ ならば、この条件は満たされる。

対照群の因果効果 TEU が DID と等しくなるためには「もし政策を実行したときの経時変化が 2 つのグループ間で等しい」ことが条件になる。

$$DID = \frac{1}{N_1} \sum_{i:z_i=1}^{N_1} (y_{bi} - y_{ai}) - \frac{1}{N_2} \sum_{i:z_i=0}^{N_2} (y_{bi} - y_{ai})$$

セミパラメトリックな”差分の差”推定

Abadie(2005)

共変量 x を考える

$$TET = E(y_{1b} - y_{0b} | z = 1) = Ex(E(y_{1b} - y_{0b} | z = 1, x)$$

$$DID = Ex(E(y_{1b} - y_{0a} | z = 1, x) - E(y_{0b} - y_a | z = 0, x))$$

$$= Ex(E(y_{1b} - y_a | z = 1, x) - E(y_{0b} - y_a | z = 1, x) + \{(E(y_{0b} - y_a | z = 1, x) - E(y_{0b} - y_a | z = 0, x)\})$$

$TET = DID$ である条件は

$$E(y_{0b} - y_a | z = 1, x) = E(y_{0b} - y_a | z = 0, x)$$

さまざまな共変量を共通にしたとき、「政策の対照にならなかったときの経時変化が 2 つのグループ間で等しい」という条件になる。

Abadie(2005) は、回帰関数を推定する代わりに傾向スコア $e = E(z - 1 | x)$ を用いて TET を表現する。

$$\begin{aligned} E(y_{rb} - y_{0b} | z = 1) &= \int E(y_{1b} - y_{0b} | z = 1, x)p(x | z = 1)dx = \\ Ex\left[E(\rho(y_b - y_a) | x)^{\frac{p(z=1|x)}{p(z=1)}}\right] &= E\left(\frac{y_b - y_a}{p(z=1)} \frac{z - e}{1 - e}\right) \\ \text{ただし } \rho &= \frac{z - e}{e(1 - e)} \text{ とすると、 } \rho = \frac{1}{e} \quad (z = 1 \text{ のとき }) \quad \rho = \frac{-1}{1 - e} \quad (z = 0 \text{ のとき }) \end{aligned}$$

$$E(\rho(y_b - y_a) | x) = E(\rho(y_b - y_a) | z = 1, x) \times e + E(\rho(y_b - y_a) | z = 1, x) - E(y_b - y_a | z = 0, x)$$

TET を傾向スコアで重み付けした推定量

$$\frac{1}{N} \sum_{i=1}^N \frac{y_{bi} - y_{ai}}{p(z=1)} \frac{z_i - e_i}{1 - e_i} \quad \text{or} \quad \frac{\sum_{i=1}^N (y_{bi} - y_{ai})^{\frac{z_i - e_i}{1 - e_i}}}{\sum_{i=1}^N e_i}$$

問題点

この推定量は、差分の差という単純な形にはなっていない。

$z = 0$ の場合の共変量の情報を有効に活用していない。

クロスセクションデータを利用した差分の差推定

b 時点の調整であれば $\delta = 1$

a 時点の調整であれば $\delta = 0$

$$\begin{array}{l} \text{処置群の調整であれば } z_a \\ \left\{ \begin{array}{l} = 1 \\ = 0 \end{array} \right. \end{array}$$

$$\begin{array}{l} \text{対照群の調整であれば } z_b \\ \left\{ \begin{array}{l} = 1 \\ = 0 \end{array} \right. \end{array}$$

$$y_b = z_b y_{1b} + (1 - z_b) y_{0b}$$

a 時点では介入がないので、 $z_a = 0, 1$ どちらでも y_a が得られる。

$$y = \delta \{ z_b y_{1b} + (1 - z_b) y_{0b} \} + (1 - \delta) y_a$$

$$\begin{aligned} DID &= E(y \mid \delta = 1, z_b = 1) - E(y \mid \delta = 0, z_a = 1) - \{E(y^- \mid \delta = 1, z_b = 0) - E(y \mid \delta = 0, z_a = 0)\} \\ &= E(y_{1b} \mid \delta = 1, z_b = 1) - E(y_a \mid \delta = 0, z_a = 1) - \{E(y_{ab} \mid \delta = 1, z_b = 0) - E(y_a \mid \delta = 0, z_a = 0)\} \end{aligned}$$

推定量

$$\frac{\sum_{i=1}^N \delta_i z_{bi} y_i}{\sum_{i=1}^N \delta_i z_{bi}} - \frac{\sum_{i=1}^N (1-\delta_i) z_{ai} y_i}{\sum_{i=1}^N (1-\delta_i) z_{ai}} - \left\{ \frac{\sum_{i=1}^N (1-z_{bi}) y_i}{\sum_{i=1}^N (1-z_{bi})} - \frac{\sum_{i=1}^N (1-\delta_i)(1-z_{ai}) y_i}{\sum_{i=1}^N (1-\delta_i)(1-z_{ai})} \right\}$$

以下の 3 条件が成立すれば、 $DID = TET$ となる。

2 時点間で調査対象者は等質である

$$E(y_{1b} \mid z_b, \delta = 1) = E(y_{1b} \mid z_b, \delta = 0)$$

$$E(y_{0b} \mid z_b, \delta = 1) = E(y_{0b} \mid z_b, \delta = 0) = E(y_{0b} \mid z_b)$$

$$E(y_a \mid z_a, \delta = 1) = E(y_a \mid z_a, \delta = 0)$$

介入しなかった場合の δ 課変数の変化が b 時点における処理群と対照群が等しい。

$$E(y_{0b} - y_a \mid z_b = 1) = E(y_{0b} - y_a \mid z_b = 0)$$

a 時点での γ 課変数の平均は 2 つのデータの処理群で共通であり、対照群でも共通。

$$E(y_a \mid z_b) = E(y_a \mid z_a)$$

$$\begin{aligned} DID &= E_x [E(y_{1b} \mid \delta = 1, z_b = 1, x) - E(y_a \mid \delta = 0, z_a = 1, x)] - \{E(y_{0b} \mid \delta = 1, z_b = 0, x) - E(y_a \mid \delta = 0, z_a = 0, x)\} \\ &\text{が } TET = E(y_{1b} - y_{0b} \mid z_b = 1) \text{ と一致する。} \end{aligned}$$

Triple differences (TD): Difference in Differences in Differences²

DD assumes that all people in region 1 at time b were treated. TD is relevant if only some qualified people ($g=1$, e.g. age) in regional at time b are treated.

²以下は、Lee(2005, pp.111-115) より引用。

$$\begin{aligned}
 g_i &= \begin{cases} 1 & \text{if treatment qualified in group1} \\ 0 & \text{if in group0} \end{cases} \\
 r_i &= \begin{cases} 1 & \text{if living in region1} \\ 0 & \text{if living in region0} \end{cases} \\
 t_t &= \begin{cases} 1 & \text{if } t = b \\ 0 & \text{if } t = a \end{cases}
 \end{aligned}
 \quad d_{it} = g_i r_i t_t \text{ for treatment}$$

y_{jit} is the potential outcome for individual i at time t , $j = 0, 1$.

$$\begin{aligned}
 E(y_{1b} - y_{0a} | g = 1, r = 1) &\\
 E(y_{0b} - y_{0a} | g = 1, r = 0) &\left. \right\} \text{first DD} \\
 E(y_{0b} - y_{0a} | g = 0, r = 1) &\\
 E(y_{0b} - y_{0a} | g = 0, r = 0) &\left. \right\} \text{second DD}
 \end{aligned}$$

DDD or TD

$$= E(y_{1b} - y_{0a} | g = 1, r = 1) - E(y_{0b} - y_{0a} | g = 1, r = 0) - \{E(y_{0b} - y_{0a} | g = 0, r = 1) - E(y_{0b} - y_{0a} | g = 0, r = 0)\}$$

$$E(y_{1b} - y_{0a} | g = 1, r = 1) - E(y_{0b} - y_{0a} | g = 1, r = 1) + \{E(y_{0b} - y_{0a} | g = 1, r = 1) - E(y_{0b} - y_{0a} | g = 1, r = 0)\} - \{E(y_{0b} - y_{0a} | g = 0, r = 1) - E(y_{0b} - y_{0a} | g = 0, r = 0)\}$$

If we assume the same time effect in $\{\cdot\}, \{\cdot\}$

then

$$TD = E(y_{1b} - y_{0b} | g = 1, r = 1)$$

The same time-effect (DD identification condition) is

$$E(y_{0b} - y_{0a} | g = 1, r = 1) - E(y_{0b} - y_{0a} | g = 1, r = 0) = 0.$$

Linear Models for TD

$$y_{ijt} = \beta_1 + \beta_g g_i + \beta_r r_i + \beta_t t_t + \beta_{rt} r_i t_t + \beta_{aj} + u_{j�}$$

$$j = 0, 1, i = 1, \dots, N, t = a, b \quad E(u_{j�}) = 0$$

$$E(u_{j�} | g, r) = 0 \quad \forall g \& r.$$

$$E(y_{1b} - y_{0a} | g = 1, r = 1) = \beta_1 + \beta_g + \beta_r + \beta_t + \beta_{rt} + \beta_d - (\beta_1 + \beta_g + \beta_r) =$$

$$\beta_t + \beta_{rt} + \beta_d$$

(Time, region*time, treatment)

$$E(y_{0b} - y_{0a} | g = 1, r = 0) = \beta_1 + \beta_g - (\beta_1 + \beta_g) = \beta_t \quad (\text{time})$$

$$E(y_{0b} - y_{0a} | g = 0, r = 0) = \beta_1 + \beta_g + \beta_t + \beta_{rt} - (\beta_1 + \beta_r) = \beta_r + \beta_{rt}$$

(time, region*time)

$$E(y_{0b} - y_{0a} | g = 0, r = 0) = \beta_1 + \beta_t - \beta_1 = \beta_t \quad (\text{time})$$

First time difference is $\beta_{rt} + \beta_d$

second time difference is β_{tt}

$$TD = \beta_d$$

General regression model

$$y_{jit} = \beta_1 + \beta_\varphi g_i + \beta_r r_i + \beta_t t_t + \beta_{gr} g_i r_i + \beta_{gt} g_i t_t + \beta_{rt} r_i t_t + \beta_d + u_{ jit}$$

Four differences in TD are

$$E(y_{1b} - y_{0a} \mid g = 1, r = 1) = \beta_1 + \beta_g + \beta_r + \beta_t + \beta_{gr} + \beta_{gt} + \beta_{rt} + \beta_d + E(u_{1b} \mid g = 1, r = 1);$$

$$\begin{aligned} E(y_{0b} - y_{0a} \mid g = 1, r = 0) &= \beta_1 + \beta_g + \beta_t + \beta_{gt} + E(u_{0b} \mid g = 1, r = 0) - \{\beta_1 + \beta_g + E(u_{0a} \mid g = 1, r = 0)\} \\ &= \beta_t + \beta_{gt} + E(u_{0b} - u_{0a} \mid g = 1, r = 0) \end{aligned}$$

$$\begin{aligned} E(y_{0b} - y_{0a} \mid g = 0, r = 1) &= \beta_1 + \beta_r + \beta_t + \beta_{rt} + E(u_{0b} \mid g = 0, r = 1) - \{\beta_1 + \beta_r + E(u_{0a} \mid g = 0, r = 1)\} \\ &= \beta_t + \beta_{rt} + E(u_{0b} - u_{0a} \mid g = 0, r = 1); \end{aligned}$$

$$\begin{aligned} E(y_{0b} - y_{0a} \mid g = 0, r = 0) &= \beta_1 + \beta_t + E(u_{0b} \mid g = 0, r = 0) - \{\beta_1 + E(u_{0a} \mid g = 0, r = 0)\} \\ &= \beta_t + E(u_{0b} - u_{0a} \mid g = 0, r = 0) \end{aligned}$$

DD in the first two differences is

$$DD_f = \beta_{rt} + \beta_d + E(u_{1b} - u_{0a} \mid g = 1, r = 1) - E(u_{0b} - u_{0a} \mid g = 1, r = 0)$$

DD in the second two differences is

$$DD_s = \beta_{rt} + E(u_{0b} - u_{0a} \mid g = 0, r = 1) - E(u_{0b} - u_{0a} \mid g = 0, r = 0)$$

Then

$$\begin{aligned} TD &= DD_f - DD_s \\ &= \beta_d + E(u \mid q - u_{0a} \mid g = 1, r = 1) - E(u_{0b} - u_{0a} \mid g = 1, r = 0) - \{E(u_{0b} - u_{0a} \mid g = 0, r = 1) - E(u_{0b} - u_{0a} \mid g = 0, r = 0)\} \end{aligned}$$

If these error terms are zero, then $TD = \beta_d$

参考文献

- [1] 星野嵩宏 (2009) 『調査観察データの統計科学』、岩波書店
- [2] 北村行伸 (2009) 『ミクロ計量経済学入門』、日本評論社
- [3] Anderson, T.W. (1984) *Introduction to Multivariate Statistical Analysis*, Wiley.

- [4] Anderson, T.W. and Rubin, H.(1949) “Estimators of the Paramters of a Single Equation in a Complete Set of Stochastic Equations”, *Annals of Mathematical Statistics*, 21, pp.570-82.
- [5] Andrew, Donald W.K. and Stock, James H.(2005) “Inference with Weal Instruments”, NBER Technical Working Paper 313.
- [6] Angrist, J.D.and Krueger, A.B.(1991) “Does Compulsory School At-tendance Affect Schooling and Earnings?”, *Quarterly Journal of Eco-nomics*, 106, pp.979-1014.
- [7] Angrist, Joshua, D., Imbens, Guid, W. and Rubin, Donald B.(1996) “Identification of Causal Effects Using Instrumental Variables”, *Jour-nal of the American Statistical Association*, 91(434), pp.444-455.
- [8] Angrist, Joshua D. and Krueger, Alan B.(2001) “Instrumental Vari-ables and the Search for Identification: From Supply and Demand to Natural Experiments”, *Journal of Economic Perspectives*, 15(4), pp.69-85.
- [9] Angrist, Joshua, D.,Bettinger, Eric., Bloom, Erik., King, Elizabeth., and Kremer, Michael. (2002) “Vouchers for Private Schooling in Colom-bia: Evidence from a Randomized Natural Experiment”, *American Economic Review*, 92(5), pp.1535-1558.
- [10] Basmann, R.L. (1960) “On Finite Sample Distributions of Generalized Classical Linear Identifiability Test Statistics”, *Journal of the Ameri-can Statistical Association*, 55(292), pp.650-59.
- [11] Baum, Christopher. (2006) *An Introduction to Modern Econometrics Using Stata*, Stata Press.
- [12] Blackburn, McKinley and Neumark, David.(1992) “Unobserved Abil-ity, Efficiency Wages, and Interindustry Wage Differentials”, *Quaterly Journal of Economics*, 107(4), pp.1421-1436.
- [13] Bound, John., Jaeger, David.A. and Baker, Regina. M.(1995) “Prob-lems with Instrumental Variables Estimation when the Correlation between the Instruments and the Endogenous Explanatory Variable is Weak”, *Journal of the American Statistical Association*, 90(430), pp.443-50.
- [14] Bowden, R.J. and Turkington, D.A.(1984) *Instrumental Variables*, Cambridge University Press.

- [15] Breusch, Trevor., Qian, Hailong., Schmidt, Peter., and Wyhowski, Donald.(1999) "Redundancy of Moment Conditions", *Journal of Econometrics*, 91, pp.89-111.
- [16] Cameron, A.C.and Trivedi, P.K.(1998) *Regression Analysis of Count Data*, Cambridge University Press.
- [17] Cameron, A.C. and Trivedi, P.K.(2005) *Microeconomics: Methods and Applications*, Cambridge University Press.
- [18] Chao, John.C. and Swanson, Norman R.(2005) "Consistent Estimation with a Large Number of Weak Instruments", *Econometrica*, 73(5), PP.1673-1692.
- [19] Cragg, John G.and Donald, Stephen G.(1993) "Testing Identifiability and Specification in Instrumental Varaible Models", *Econometric Theory*, 9, pp.222-40.
- [20] Davidson, Russell and MacKinnon, James G.(2004) *Econometric Theory and Methods*, Oxford University Press.
- [21] Durbin, J.(1954) "Errors in variables", *Review of the Internatinal Statistical Institute*, 22, pp.23-32.
- [22] Griliches, Zvi.(1976) "Wages of Very Young Men", *Journal of Political Economy*, 84(4. Part 2), pp. S69-S85.
- [23] Griliches, Zvi.(1977) "Estimating the Returns to Schooling: Some Econometric Problems", *Econometrica*, 45(1), pp.1-22.
- [24] Griliches, Zvi., Hall, Bronwyn., and Hausoman, Jerry.(1978) "Missing Data and self-Selection in Large Panels", *Annales de L'INSEE*, XXX-XXXI, pp.137-76.
- [25] Hahn, Jinyoung and Hausman, Jerry. (2002a) "A New Specification Test for the Validity of Instrumental variables", *Econometrica*, 70(1), pp.163-189.
- [26] Hahn, Jinyoung and Hausman, Jerry. (2002b) "Notes on Bias in Estimators for Simultaneous Equation Models", *Economics Letters*, 75. pp.237-241.
- [27] Hahn, Jinyoung and Hausman, Jerry. (2003) "Weak Instruments: Diagnosis and Cures in Empirical Econometrics", *American Economic Review*, 93(2), pp.118-125.

- [28] Hall, Alastair R., Rudebusch, Glenn D. and Wilcox, David W.(1996) “Judging Instrument Relevance in Instrumental Variables Estimation”, *International Economic Review*, 37(2), pp.283-298.
- [29] Hall, Alastair R. and Peixe, Fernanda P.M.(2000) “A Consistent Method for the Selection of Relevant Instruments”, A paper presented at Econometric Society World Congress 2000.
- [30] Hansen, Lars.P (1982) “Large Sample Properties of Generalized Method of Moments Estimators”, *Econometrica*, 50(4), pp.1029-1054.
- [31] Hausman, Jerry. (1978) “Specification tests in econometrics”, *Econometrica*, 46, pp.1251-72.
- [32] Hausman, Jerry., Stock, James H. and Yogo, Motohiro.(2005) “Asymptotic Properties of the Hahn-Hausman Test for Weak-Instruments”, *Economics Letters*, 89, pp.333-42.
- [33] Hayashi, Fumio.(2000) *Econometrics*, Princeton University Press.
- [34] Imbens, Guido W. and Angrist, Joshua D.(1994) “Identification and Estimation of Local Average Treatment Effects”, *Econometrica*, 62(2). pp.467-475.
- [35] Koenker, Roger.(1981) “A Note on Studentizing a test for Heteroscedasticity”, *Journal of Econometrics*, 17., pp.107-112.
- [36] Koenker, Roger. (2005) *Quantile Regression*, Cambridge University Press.
- [37] Nelson, Charles R. and Startz, Richard.(1990a) “The Distribution of the Instrumental Variables Estimator and Its t-Ratio When the Instrument is a Poor One”, *Journal of Business*, 63(1, Part.2), pp. S125-S140.
- [38] Nelson, Charles R.and Startz, Richard.(1990) “Some Further Results on the Exact Small Sample Properties of the Instrumental Variable Estimator”, *Econometrica*, 58(4), pp.967-76.
- [39] Pagan, A.R. and Hall, D. (1983) “Diagnostic Tests as Residual Analysis”, *Econometric Reviews*, 2(2), pp.159-218.
- [40] Ruud, P.A. (2000) *An Introduction to Classical Econometric Theory*, Oxford University Press.
- [41] Sargan, J.D. (1958) “The Estimation of Economic Relationships Using Instrumental Variables”, *Econometrica*, 26(3), pp.393-415.

- [42] Shea, John.(1997) "Instrument Relevance in Multivariate Linear Models: A Simple Measure", *Review of Economics and Statistics*, 79(2), pp.348-352.
- [43] Staiger, Douglas. and Stock, James.H. (1997) "Instrumental Variables Regression with Weak Instruments", *Econometrica*, 65(3), pp.557-86.
- [44] Stock, James H. and Wright Jonathan H. (2000) "GMM with Weak Identification", *Econometrica*, 68(5), pp.1055-96.
- [45] Stock, James H., Wright, Jonathan H. and Yogo, Motohiro. (2002) "A Survey of Weak Instruments and Weak Identification in Generalized Method of Moments", *Journal of Business and Economic Statistics*, 20(4), pp.518-29.
- [46] Stock, James H. and Yogo, Motohiro. (2005) "Testing for Weak Instruments in Linear IV Regression", in Andrews, D.W.K. and Stock, J.H.(eds) *Identification and Inference for Econometric Models: Essays in Honor of Thomas Rothenberg*, Cambridge University Press. pp.80-108.
- [47] Winkelmann, Rainer and Boes, Stefan. (2005) *Analysis of Microdata*, Springer.
- [48] White, Halbert. (1980) "A Heteroskedasticity-Consistent Covariance Matrix Estimator and a Direct Test for Heteroskedasticity", *Econometrica*, 48(4), pp.817-838.
- [49] White, Halbert. (1982) "Instrumental Variables Regression with Independent Observations", *Econometrica*, 50(2), pp.483-499.
- [50] Wooldridge, Jeffrey. M. (2002) *Econometric Analysis of Cross Section and Panel Data*, The MIT Press
- [51] Wu, D-M. (1973) "Alternative tests of independence between stochastic regressors and disturbances", *Econometrica*, 41, pp.733-50.